Learning linear operators: Infinite-dimensional regression as a well-behaved non-compact inverse problem

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For all the details:



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- Let X and Y be Bochner square-integrable random variables taking values in separable Hilbert spaces \mathcal{X} and \mathcal{Y} respectively, i.e. $(X, Y) \in L^2(\mathbb{P}; \mathcal{X} \times \mathcal{Y})$.
- We aim to solve the following regression problem:

minimise
$$\mathbb{E}[\|Y - \theta X\|_{\mathcal{Y}}^2] \equiv \|Y - \theta X\|_{L^2(\mathbb{P};\mathcal{Y})}^2$$
 w.r.t. $\theta \in L(\mathcal{X},\mathcal{Y}),$ (RP)

where $L(\mathcal{X}, \mathcal{Y})$ is the Banach space of bounded linear operators from \mathcal{X} into \mathcal{Y} .

- In practice, we will only have data points (X_i, Y_i), i = 1,..., n so we must think about empirical approximation and regularisation.
- Moral of the talk: From a regularisation standpoint, (RP) is "just as hard" as finite-dimensional regression in reasonable settings.

- If at least one of \mathcal{X} and \mathcal{Y} has infinite dimension, then so too does the search space $L(\mathcal{X}, \mathcal{Y})$, and so (RP) is an infinite-dimensional regression problem.
- We are particularly motivated by the case of infinite-dimensional \mathcal{Y} , exemplified by relevant applications in
 - functional linear regression with functional response (Ramsay and Silverman, 2005);
 - non-parametric regression with vector-valued kernels (Caponnetto and De Vito, 2007) (more on this in a moment);
 - the conditional mean embedding (Park and Muandet, 2020; Li et al., 2022);
 - and inference for Hilbertian time series (Bosq, 2000).

Example: (Vector-valued) kernel regression 1

- Let \mathcal{E} be a second-countable locally compact Hausdorff space equipped with its Borel σ -algebra $\mathcal{B}_{\mathcal{E}}$, and let \mathcal{X} be an RKHS of \mathbb{R} -valued functions on \mathcal{E} with reproducing kernel $k: \mathcal{E}^2 \to \mathbb{R}$ and canonical feature map $\varphi: \mathcal{E} \to \mathcal{X}$.
- Assume further that (*E*, *B_E*) is equipped with a probability measure μ, with a compact embedding operator i: *X* → *L*²(μ) (e.g. Christmann and Steinwart, 2008, Section 4.3).
- Let 𝒴 be another separable real Hilbert space. Consider 𝔅 := {Aφ(·) | A ∈ S₂(𝔅, 𝔅)}; this is a vv-RKHS of 𝒴-valued functions with operator-valued reproducing kernel

 $\mathcal{K} \colon \mathcal{E}^2 \to \mathcal{L}(\mathcal{Y})$ $(x, x') \mapsto k(x, x') \operatorname{Id}_{\mathcal{Y}}$

and we have a bounded linear embedding operator

$$I := i \otimes \mathrm{Id}_{\mathcal{Y}} \colon \mathcal{G} \cong \mathcal{X} \otimes \mathcal{Y} \hookrightarrow L^{2}(\mu) \otimes \mathcal{Y} \cong L^{2}(\mu; \mathcal{Y}).$$

As the embedding $i: \mathcal{X} \hookrightarrow L^2(\mu)$ is compact, the embedding $I := i \otimes Id_{\mathcal{Y}}$ is compact $\iff \dim \mathcal{Y} < \infty$.

- We now consider an *E*-valued random variable ξ with law *L*(ξ) =: μ on (*E*, *B_E*) and a *Y*-valued random variable *Y*, both defined on a common probability space.
- The nonlinear kernel regression problem

$$\min_{F\in\mathcal{G}}\mathbb{E}[\|Y-F(\xi)\|_{\mathcal{Y}}^2]$$

is equivalent to the (Hilbert–Schmidt) version of the linear regression problem (RP) with $X := \varphi(\xi)$:

$$\min_{\theta \in S_2(\mathcal{X}, \mathcal{Y})} \mathbb{E}[\|Y - \theta \varphi(\xi)\|_{\mathcal{Y}}^2].$$

Problem reformulation

The problem with infinite-dimensional regression

- Infinite-dimensional linear regression does not necessarily admit a minimiser!
- Assuming a well-specified linear model, i.e. the existence of a bounded linear operator $\theta_{\star} \colon \mathcal{X} \to \mathcal{Y}$ such that

$$Y = \theta_{\star}X + \varepsilon$$

with an exogeneous \mathcal{Y} -valued noise variable ε satisfying $\mathbb{E}[\varepsilon|X] = 0$, (RP) is equivalent to the operator factorisation problem

$$C_{YX} = \theta C_{XX}, \qquad \theta \in L(\mathcal{X}, \mathcal{Y}),$$
 (OFP)

where $C_{YX} \in L(\mathcal{X}, \mathcal{Y})$ and $C_{XX} \in L(\mathcal{X}, \mathcal{X})$ are the covariance operators (Baker, 1973) associated with X and Y.

 Solubility of (OFP) is related to a well-known set of range inclusion and operator majorisation conditions due to Douglas (1966) and the Moore-Penrose pseudoinverse (Engl et al., 1996).

Recap: Tensor products and covariance operators

• For $y \in \mathcal{Y}$ and $x \in \mathcal{X}$, $y \otimes x \in L(\mathcal{X}, \mathcal{Y})$ is the rank-one operator

 $\mathcal{X} \ni \mathbf{v} \mapsto (\mathbf{y} \otimes \mathbf{x})(\mathbf{v}) \coloneqq \langle \mathbf{x}, \mathbf{v} \rangle_{\mathcal{X}} \mathbf{y} \in \mathcal{Y}.$

- The Hilbert tensor product 𝒴 ⊗ 𝒴 is defined to be the completion of the linear span of all such rank-one operators w.r.t. ⟨𝒴 ⊗ 𝑥, 𝒴' ⊗ 𝑥'⟩𝒴 ⊗𝑥 := ⟨𝒴, 𝒴'⟩𝒴 ⟨𝑥, 𝑥'⟩𝑥.
- Note that 𝒴 ⊗ 𝒴 is isometric with S₂(𝒴, 𝒴), the space of Hilbert–Schmidt operators; and also L²(ℙ; 𝒴) ≅ L²(ℙ; ℝ) ⊗ 𝒴.
- The (uncentred) covariance operators (Baker, 1973) of Y with X, and of X with itself, are given by

$$\mathbb{C}ov[Y, X] \coloneqq C_{YX} \coloneqq \mathbb{E}[Y \otimes X] \in S_1(\mathcal{X}, \mathcal{Y}) = \{ \text{trace-class op's} \} \text{ and } \\ \mathbb{C}ov[X, X] \coloneqq C_{XX} \coloneqq \mathbb{E}[X \otimes X] \in S_1(\mathcal{X}).$$

- Note that $C_{YX}^* = C_{XY}$, and so C_{XX} is self-adjoint.
- The covariance operators are the unique operators satisfying

$$\mathbb{E}[\langle y, Y \rangle_{\mathcal{Y}} \langle x, X \rangle_{\mathcal{X}}] = \langle y, C_{YX} x \rangle_{\mathcal{Y}} \quad \text{for all } x \in \mathcal{X}, y \in \mathcal{Y}.$$

The operator factorisation problem (OFP)

$$C_{YX} = \theta C_{XX}, \qquad \theta \in L(\mathcal{X}, \mathcal{Y}),$$
 (OFP)

can be reformulated in terms of a (potentially ill-posed) linear inverse problem

$$A_{C_{XX}}[\theta] = C_{YX}, \quad \theta \in L(\mathcal{X}, \mathcal{Y})$$
(IP)

based on the (generally non-compact) forward operator $A_{C_{XX}}$: $L(\mathcal{X}, \mathcal{Y}) \rightarrow L(\mathcal{X}, \mathcal{Y})$,

$$A_{C_{XX}}[\theta] \coloneqq \theta C_{XX}.$$

- We call the operator $A_{C_{XX}}$ the precomposition operator associated with C_{XX} .
- Even in the misspecified case, the solution to the inverse problem (IP) still characterises the minimiser of the linear regression problem (RP)!

Spectral theory and regularisation

• The standard, naïve thing to do at this point would be to solve (IP)

$$A_{C_{XX}}[\theta] = C_{YX}, \quad \theta \in L(\mathcal{X}, \mathcal{Y})$$

using the Moore–Penrose pseudoinverse of $A_{C_{XX}}$:

$$\theta = \mathbf{A}^{\dagger}_{\mathbf{C}_{\mathbf{X}\mathbf{X}}}[\mathbf{C}_{\mathbf{Y}\mathbf{X}}].$$

- The problem is that dim $\mathcal{Y} = \infty \implies A_{C_{XX}}$ is non-compact, in which case we have no good off-the-shelf spectral theory for $A_{C_{XX}}$, no pseudoinverse, etc.
- Fortunately, we can build a decent spectral theory for A_{CXX} if we focus on the Hilbert–Schmidt setting: we restrict the search to θ ∈ S₂(X, Y) and use the fact that

$$A_{\mathcal{C}_{XX}} \colon S_2(\mathcal{X}, \mathcal{Y}) \to S_2(\mathcal{X}, \mathcal{Y}).$$

Theorem 1 (Spectral decomposition)

Let $C \in S_2(\mathcal{X})$ be self-adjoint with spectral decomposition

$$\mathcal{C} = \sum_{\lambda \in \sigma_{\mathsf{p}}(\mathcal{C})} \lambda \mathcal{P}_{\mathsf{eig}_{\lambda}(\mathcal{C})} \,,$$

where $P_{eig_{\lambda}(C)}: \mathcal{X} \to \mathcal{X}$ is orthogonal projection onto $eig_{\lambda}(C)$ and the above series expression converges in operator norm. Then the non-compact induced precomposition operator A_C on $S_2(\mathcal{X}, \mathcal{Y})$ has pure point spectrum and the spectral decomposition

$$A_{\mathcal{C}} = \sum_{\lambda \in \sigma_{\mathsf{p}}(\mathcal{C})} \lambda P_{\mathcal{Y} \otimes \mathsf{eig}_{\lambda}(\mathcal{C})},$$

where $P_{\mathcal{Y} \otimes eig_{\lambda}(C)}$: $S_2(\mathcal{X}, \mathcal{Y}) \to S_2(\mathcal{X}, \mathcal{Y})$ is orthogonal projection onto $\mathcal{Y} \otimes eig_{\lambda}(C)$ and the above series converges in operator norm.

Spectral theory for precomposition operators 2

Corollary 2 (Compatibility with functional calculus)

Let $C = \sum_{\lambda \in \sigma_p(C)} \lambda P_{eig_{\lambda}(C)} \in S_2(\mathcal{X})$ be self-adjoint. If $g : \mathbb{R} \to \mathbb{R}$ is extended to act on self-adjoint Hilbert space operators with pure point spectrum in terms of their spectral decompositions via

$$g(C) \coloneqq \sum_{\lambda \in \sigma_{\mathsf{p}}(C)} g(\lambda) P_{\mathsf{eig}_{\lambda}(C)},$$

then A_C as an operator on $S_2(\mathcal{X}, \mathcal{Y})$ satisfies

$$A_{g(C)} = g(A_C) = \sum_{\lambda \in \sigma_p(C)} g(\lambda) P_{\mathcal{Y} \otimes eig_{\lambda}(C)},$$

We will use this with $g = g_{\alpha}$ being some approximation — e.g. Tikhonov, spectral cutoff, ... — to the 'ideal' inverse $g(\lambda) = \lambda^{-1}$, yielding a regularised population solution to (IP):

$$\theta_{\alpha} \coloneqq g_{\alpha}(A_{C_{XX}})[C_{YX}] = C_{YX}g_{\alpha}(C_{XX}).$$

Terminology for regularisation

A family of functions g_α: [0,∞) → ℝ, indexed by a regularisation parameter α > 0, is a spectral regularisation strategy (Engl et al., 1996) if

(R1)
$$\sup_{\lambda \in [0,\infty)} |\lambda g_{\alpha}(\lambda)| \leq D$$
 for some constant D ,

(R2) $\sup_{\lambda \in [0,\infty)} |1 - \lambda g_{\alpha}(\lambda)| \leqslant \gamma_0$ for some constant γ_0 , and

(R3) $\sup_{\lambda \in [0,\infty)} |g_{\alpha}(\lambda)| < B\alpha^{-1}$, for some constant *B*.

- We write $r_{\alpha}(\lambda) := 1 \lambda g_{\alpha}(\lambda)$ for the residual associated to the regularisation scheme g_{α} .
- The qualification of g_{α} is the maximal q such that

$$\sup_{\lambda \in [0,\infty)} \lambda^q |r_\alpha(\lambda)| \equiv \sup_{\lambda \in [0,\infty)} \lambda^q |1 - \lambda g_\alpha(\lambda)| \leqslant \gamma_q \alpha^q$$

for some constant γ_q which does not depend on α .

 Such assumptions are also common in learning theory (see e.g. Bauer et al., 2007; Gerfo et al., 2008; Dicker et al., 2017; Blanchard and Mücke, 2018). **Regularised empirical solutions**

Empirical solutions

- X and Y are in practice only accessible through sample pairs $(X_i, Y_i) \in \mathcal{X} \times \mathcal{Y}$ for i = 1, ..., n.
- For simplicity, we assume that these sample pairs are obtained i.i.d. from the joint law of (X, Y).
- We define the empirical covariance operators by

$$\widehat{C}_{XX} \coloneqq \frac{1}{n} \sum_{i=1}^{n} X_i \otimes X_i \text{ and } \widehat{C}_{YX} \coloneqq \frac{1}{n} \sum_{i=1}^{n} Y_i \otimes X_i.$$

- Note that \widehat{C}_{XX} and \widehat{C}_{XY} are \mathbb{P} -a.s. of rank at most n.
- We now analyse the regularised empirical solution

$$\widehat{\theta}_{\alpha} \coloneqq g_{\alpha}(A_{\widehat{\mathcal{C}}_{XX}})[\widehat{\mathcal{C}}_{YX}] = \widehat{\mathcal{C}}_{YX} g_{\alpha}(\widehat{\mathcal{C}}_{XX}). \tag{EMP}$$

- We can obtain rates for Hilbert-Schmidt regression based on Hölder source conditions.
- We analyse the error $\theta_{\star} \hat{\theta}_{\alpha}$ associated with the regularised empirical solution $\hat{\theta}_{\alpha}$.
- In particular, we are interested both in the Hilbert–Schmidt norm of this error and in the mean-square prediction error

$$\mathbb{E}\left[\left\|\left(\theta_{\star}-\widehat{\theta}_{\alpha}\right)X\right\|_{\mathcal{Y}}^{2}\right]\equiv\left\|\left(\theta_{\star}-\widehat{\theta}_{\alpha}\right)C_{XX}^{1/2}\right\|_{\mathcal{S}_{2}(\mathcal{X},\mathcal{Y})}^{2}.$$

• To treat these in a unified way we will examine

$$\left\| \left(\theta_{\star} - \widehat{\theta}_{\alpha} \right) C^{s}_{XX} \right\|_{\mathcal{S}_{2}(\mathcal{X}, \mathcal{Y})} \text{ for } 0 \leqslant s \leqslant \frac{1}{2}.$$

To establish quantitative convergence rates, we need a priori assumptions on the "smoothness" of the ground truth θ_{\star} , a.k.a. "source conditions":

Assumption 3

We assume that the solution satisfies the Hölder source condition $\theta_{\star} \in \Omega(\nu, R)$, where

$$\Omega(\nu, R) := \left\{ A_{\mathcal{C}_{XX}}^{\nu}[\theta] \, \big| \, \theta \in \mathcal{S}_{2}(\mathcal{X}, \mathcal{Y}), \|\theta\|_{\mathcal{S}_{2}(\mathcal{X}, \mathcal{Y})} \leqslant R \right\} \subseteq \mathcal{S}_{2}(\mathcal{X}, \mathcal{Y}).$$

Lemma 4

The source condition $\theta_{\star} \in \Omega(\nu, R)$ holds if and only if the moment condition

$$\sum_{i \in I} \sup_{x \in \mathcal{X}} \frac{\left| \mathbb{E}[\langle x, X \rangle_{\mathcal{X}} \langle e_i, Y \rangle_{\mathcal{Y}}] \right|^2}{\| C_{XX}^{\nu+1} x \|_{\mathcal{X}}^2} \leqslant R^2$$

hold for some (indeed, any) complete orthonormal system $\{e_i\}_{i \in I}$ in \mathcal{Y} .

Decomposing the error 1/2

• Naïve error decomposition: $\mathbb{P}^{\otimes n}$ -a.s. with respect to the samples $(X_i, Y_i)_{i=1}^n$,

$$\left\| \left(\theta_{\star} - \widehat{\theta}_{\alpha} \right) C_{XX}^{s} \right\|_{S_{2}(\mathcal{X}, \mathcal{Y})} \leqslant \underbrace{\left\| \left(\theta_{\star} - \theta_{\alpha} \right) C_{XX}^{s} \right\|_{S_{2}(\mathcal{X}, \mathcal{Y})}}_{= \text{approximation error}} + \underbrace{\left\| \left(\theta_{\alpha} - \widehat{\theta}_{\alpha} \right) C_{XX}^{s} \right\|_{S_{2}(\mathcal{X}, \mathcal{Y})}}_{= \text{variance}} \right).$$
(3.1)

However, this decomposition turns out to be less than ideal and instead we use:

$$\begin{aligned} \theta_{\star} - \widehat{\theta}_{\alpha} &= \theta_{\star} - \theta_{\star} \widehat{C}_{XX} g_{\alpha} (\widehat{C}_{XX}) + \theta_{\star} \widehat{C}_{XX} g_{\alpha} (\widehat{C}_{XX}) - \widehat{\theta}_{\alpha} \\ &= \theta_{\star} r_{\alpha} (\widehat{C}_{XX}) + \theta_{\star} \widehat{C}_{XX} g_{\alpha} (\widehat{C}_{XX}) - \widehat{C}_{YX} g_{\alpha} (\widehat{C}_{XX}) \\ &= \theta_{\star} r_{\alpha} (\widehat{C}_{XX}) + (\theta_{\star} \widehat{C}_{XX} - \widehat{C}_{YX}) g_{\alpha} (\widehat{C}_{XX}). \end{aligned}$$

• Hence, $\mathbb{P}^{\otimes n}$ -a.s.,

$$\left\| \left(\theta_{\star} - \widehat{\theta}_{\alpha}\right) C_{XX}^{s} \right\|_{S_{2}(\mathcal{X},\mathcal{Y})} \leqslant \left\| \theta_{\star} r_{\alpha}(\widehat{C}_{XX}) C_{XX}^{s} \right\|_{S_{2}(\mathcal{X},\mathcal{Y})} + \left\| \left(\theta_{\star} \widehat{C}_{XX} - \widehat{C}_{YX}\right) g_{\alpha}(\widehat{C}_{XX}) C_{XX}^{s} \right\|_{S_{2}(\mathcal{X},\mathcal{Y})}.$$

• Hence, $\mathbb{P}^{\otimes n}$ -a.s.,

$$\| (\theta_{\star} - \widehat{\theta}_{\alpha}) C_{XX}^{\mathfrak{s}} \|_{S_{2}(\mathcal{X}, \mathcal{Y})} \leq \| \theta_{\star} r_{\alpha} (\widehat{C}_{XX}) C_{XX}^{\mathfrak{s}} \|_{S_{2}(\mathcal{X}, \mathcal{Y})} + \| (\theta_{\star} \widehat{C}_{XX} - \widehat{C}_{YX}) g_{\alpha} (\widehat{C}_{XX}) C_{XX}^{\mathfrak{s}} \|_{S_{2}(\mathcal{X}, \mathcal{Y})}.$$
(3.2)

- Again, we think of the two terms on the right-hand side of (3.2) as an approximation error and a variance term.
- Crucially, though, the approximation error in the decomposition (3.2) is random as opposed to the deterministic approximation term in (3.1) — and both terms in (3.2) will be amenable to analysis using concentration-of-measure techniques.

The key tool for us is a recent concentration inequality for Hilbert space-valued random variables:

Theorem 5 (Maurer and Pontil, 2021, Prop. 7.11)

Let $\xi, \xi_1, \ldots, \xi_n$ be i.i.d. random variables with joint law $\mathbb{P}^{\otimes n}$ taking values in a separable Hilbert space \mathcal{H} such that $\mathbb{E}[\xi] = 0$ and the subexponential norm $\|\xi\|_{L_{\psi_1}(\mathbb{P};\mathcal{H})}$ is finite. Then, for all $\delta \in (0, \frac{1}{2}]$ and $n \ge \log(1/\delta)$, with $\mathbb{P}^{\otimes n}$ -probability at least $1 - \delta$,

$$\left\|\frac{1}{n}\sum_{i=1}^{n}\xi_{i}\right\|_{\mathcal{H}} \leqslant 8\sqrt{2}e\|\xi\|_{L_{\psi_{1}}(\mathbb{P};\mathcal{H})}\sqrt{\frac{\log(1/\delta)}{n}}.$$

Despite the large number of terms that we need to bound, we carefully reduce the number of independent appeals to Maurer and Pontil (2021) to a minimum of only two.

Subexponential and sub-Gaussian norms

For a real-valued random variable ξ defined on (Ω, F, P), we introduce the Banach spaces L_{ψ1}(Ω, F, P; R) = L_{ψ1}(P) and L_{ψ2}(Ω, F, P; R) = L_{ψ2}(P) via the norms

$$\begin{split} \text{subexponential:} & \|\xi\|_{L_{\psi_1}(\mathbb{P})} \coloneqq \sup_{1 \leqslant p < \infty} \frac{\|\xi\|_{L^p(\mathbb{P})}}{p}, \\ \text{sub-Gaussian:} & \|\xi\|_{L_{\psi_2}(\mathbb{P})} \coloneqq \sup_{1 \leqslant p < \infty} \frac{\|\xi\|_{L^p(\mathbb{P})}}{p^{1/2}}. \end{split}$$

• For ξ taking values in a separable Hilbert space \mathcal{H} :

$$\|\xi\|_{L_{\psi_1}(\mathbb{P};\mathcal{H})} \coloneqq \|\|\xi\|_{\mathcal{H}}\|_{L_{\psi_1}(\mathbb{P})} = \sup_{1 \le p < \infty} \frac{\|\xi\|_{L^p(\mathbb{P};\mathcal{H})}}{p}$$

and analogously for $\|\xi\|_{L_{\psi_2}(\mathbb{P};\mathcal{H})} \coloneqq \|\|\xi\|_{\mathcal{H}}\|_{L_{\psi_2}(\mathbb{P})}.$

Convergence rates

Theorem 6 (Convergence rates under Hölder source conditions)

Suppose that g_{α} has qualification $q \ge \nu + s$. Suppose that $Y \in L_{\psi_2}(\mathbb{P}; \mathcal{Y})$, $X \in L_{\psi_2}(\mathbb{P}; \mathcal{X})$, $\theta_{\star} \in \Omega(\nu, R)$, and $0 < \alpha < 1$. Let $\delta \in (0, \frac{1}{e}]$ and $s \in [0, \frac{1}{2}]$. For the regularisation schedule

$$\alpha_n \coloneqq \left(\frac{1}{\sqrt{n}}\right)^{\frac{1}{\nu+1}}$$

and for

$$n \ge n_0 := \max \left\{ \|X\|_{L_{\psi_2}(\mathbb{P};\mathcal{X})}^4, \left(1152e^2 \|X\|_{L_{\psi_2}(\mathbb{P};\mathcal{X})}^4 \log(1/\delta)\right)^{\frac{1}{\nu}} \right\}^{1+\nu}$$

with $\mathbb{P}^{\otimes n}$ -probability at least $1 - 2\delta$,

$$\left\| \left(\theta_{\star} - \widehat{\theta}_{\alpha_n} \right) C_{XX}^{s} \right\|_{\mathcal{S}_2(\mathcal{X}, \mathcal{Y})} \leqslant 3\bar{\kappa} \sqrt{\log(1/\delta)} \left(\frac{1}{\sqrt{n}} \right)^{\frac{s+\nu}{1+\nu}},$$

where $\bar{\kappa}$ is an explicit constant depending only on the regularisation scheme, the source condition, and the sub-Gaussian norms of X and Y.

Optimal rates and comparison to kernel setting

- The rates in Theorem 6 match those of kernel regression with scalar and finite-dimensional response variables under a Hölder source condition and with no additional assumptions on the eigenvalue decay of C_{XX} (Caponnetto and De Vito, 2007; Blanchard and Mücke, 2018; Lin et al., 2020).
- Minimax optimality of these rates is only derived by Caponnetto and De Vito (2007) and Blanchard and Mücke (2018) under the additional assumption that the eigenvalues of C_{XX} decay rapidly enough, which is an implicit assumption on the marginal distribution of X.
- To establish minimax optimality in our setting, we would have to repeat the standard arguments, e.g. apply a general reduction scheme in conjunction with Fano's method (Tsybakov, 2009).
- However, as discussed earlier, the Hilbert–Schmidt regression problem has scalar response kernel regression and some settings of kernel regression with vector-valued response as special cases.

Closing remarks

- Can we obtain fast 1/n rates? This would require additional assumptions about the joint law of (X, Y). So far, this is only solved for the special case of the CME (Li et al., 2022).
- Solving (RP)/(IP) over the non-reflexive Banach space L(X, Y) a simple yet really evil example is X = Y and θ_⋆ = Id.
- Learning in L(X, Y) requires more general source conditions, since the Hölder source condition θ_{*} = θ̃ C^ν_{XX} with θ̃ ∈ L(X, Y) already implies the Hilbert–Schmidt setting θ_{*} ∈ S₂(X, Y) for ν > 1/2.
- For Banach space \mathcal{X} and \mathcal{Y} , a suitable analogue of (IP) is needed. The Hilbert case uses $\operatorname{tr}(C_{XX}) = \mathbb{E}[\|X\|_{\mathcal{X}}^2]$ and derivative of squared norm.
- Extension to more general non-i.i.d. sample data, e.g. autoregression for stationary time series?

Thank You!



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